

## Lecture 27

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## 1 Space-Bounded Derandomization

We now discuss derandomization of space-bounded algorithms. Here non-trivial results can be shown *without making any unproven assumptions*, in contrast to what is currently known for derandomizing time-bounded algorithms. We show first that  $\mathcal{BPL} \subseteq \text{SPACE}(\log^2 n)$  and then improve the analysis and show that<sup>1</sup>  $\mathcal{BPL} \subseteq \text{TIME}(\text{poly}(n), \log^2 n) \subseteq \mathcal{SC}$ . (Note: we already know

$$\mathcal{RL} \subseteq \text{NL} \subseteq \text{SPACE}(\log^2 n)$$

but this does not by itself imply  $\mathcal{BPL} \subseteq \text{SPACE}(\log^2 n)$ .)

With regard to the first result, we actually prove something more general:

**Theorem 1** *Any randomized algorithm (with two-sided error) that uses space  $S = \Omega(\log n)$  and  $R$  random bits can be converted to one that uses space  $\mathcal{O}(S \log R)$  and  $\mathcal{O}(S \log R)$  random bits.*

Since any algorithm using space  $S$  uses time at most  $2^S$  (by our convention regarding probabilistic machines) and hence at most this many random bits, the following is an immediate corollary:

**Corollary 2** *For  $S = \Omega(\log n)$  it holds that  $\text{BSPACE}(S) \subseteq \text{SPACE}(S^2)$ .*

**Proof** Let  $L \in \text{BSPACE}(S)$ . Theorem 1 shows that  $L$  can be decided by a probabilistic machine with two-sided error using  $\mathcal{O}(S^2)$  space and  $\mathcal{O}(S^2)$  random bits. Enumerating over all random bits and taking majority, we obtain a deterministic algorithm that uses  $\mathcal{O}(S^2)$  space. ■

## 2 $\mathcal{BPL} \subseteq \text{SPACE}(\log^2 n)$

We now prove Theorem 1. Let  $M$  be a probabilistic machine running in space  $S$  (and time  $2^S$ ), using  $R \leq 2^S$  random bits, and deciding a language  $L$  with two-sided error. (Note that  $S, R$  are functions of the input length  $n$ , and the theorem requires  $S = \Omega(\log n)$ .) We will assume without loss of generality that  $M$  always uses exactly  $R$  random bits on all inputs; recall also that  $M$  has *read-once* access to its random bits. Fixing an input  $x$  and letting  $\ell$  be some parameter, we will view the computation of  $M_x$  as a random walk on a multi-graph in the following way: our graph will have  $R/\ell + 1$  layers, with each layer containing  $N \stackrel{\text{def}}{=} 2^{\mathcal{O}(S)}$  nodes that correspond to possible configurations of  $M_x$ . There is an edge from node  $a$  (in some layer  $i$ ) to node  $b$  (in some layer  $i + 1$ ) labeled by the string  $r \in \{0, 1\}^\ell$  iff  $M_x$  moves from configuration  $a$  to configuration  $b$  after reading  $r$  as its next  $\ell$  random bits. Computation of  $M(x)$  is then equivalent to a random walk of length  $R/\ell$  on this graph, beginning from the node corresponding to the initial configuration of  $M_x$  (in

<sup>1</sup> $\mathcal{SC}$  captures computation that *simultaneously* uses polynomial time and polylogarithmic space.

layer 0). If  $x \in L$  then the probability that this random walk ends up in the accepting state is at least  $2/3$ , while if  $x \notin L$  then the probability that this random walk ends up in the accepting state is at most  $1/3$ .

It will be convenient to represent this process using an  $N \times N$  transition matrix  $Q_x$ , where the entry in column  $i$ , row  $j$  is the probability that  $M_x$  moves from configuration  $i$  to configuration  $j$  after reading  $\ell$  random bits. Vectors of length  $N$  whose entries are non-negative and sum to 1 correspond to probability distributions over the configurations of  $M_x$  in the natural way. If we let  $\mathbf{s}$  denote the probability distribution that places probability 1 on the initial configuration of  $M_x$  (and 0 elsewhere), then  $Q_x^{R/\ell} \cdot \mathbf{s}$  corresponds to the probability distribution over the final configuration of  $M_x$ ; thus, if we let  $i$  denote the accepting configuration of  $M_x$ :

$$\begin{aligned} x \in L &\Rightarrow \left(Q_x^{R/\ell} \cdot \mathbf{s}\right)_i \geq 3/4 \\ x \notin L &\Rightarrow \left(Q_x^{R/\ell} \cdot \mathbf{s}\right)_i \leq 1/4. \end{aligned}$$

The statistical difference between two vectors/probability distributions  $\mathbf{s}, \mathbf{s}'$  is

$$\text{SD}(\mathbf{s}, \mathbf{s}') \stackrel{\text{def}}{=} \frac{1}{2} \cdot \|\mathbf{s} - \mathbf{s}'\|_1 = \frac{1}{2} \cdot \sum_i |\mathbf{s}_i - \mathbf{s}'_i|.$$

If  $Q, Q'$  are two transition matrices — meaning that all entries are non-negative, and the entries in each column sum to 1 — then we abuse notation and define

$$\text{SD}(Q, Q') \stackrel{\text{def}}{=} \max_{\mathbf{s}} \{\text{SD}(Q\mathbf{s}, Q'\mathbf{s})\},$$

where the maximum is taken over all  $\mathbf{s}$  that correspond to probability distributions. Note that if  $Q, Q'$  are  $N \times N$  transition matrices and  $\max_{i,j} \{|Q_{i,j} - Q'_{i,j}|\} \leq \varepsilon$ , then  $\text{SD}(Q, Q') \leq N\varepsilon/2$ .

## 2.1 A Useful Lemma

The pseudorandom generator we construct will use a family  $H$  of pairwise-independent functions as a building block. It is easy to construct such a family  $H$  whose functions map  $\ell$ -bit strings to  $\ell$ -bit strings and such that (1)  $|H| = 2^{2\ell}$  (and so choosing a random member of  $H$  is equivalent to choosing a random  $2\ell$ -bit string) and (2) functions in  $H$  can be evaluated in  $\mathcal{O}(\ell)$  space.

For  $S \subseteq \{0, 1\}^\ell$ , define  $\rho(S) \stackrel{\text{def}}{=} |S|/2^\ell$ . We define a useful property and then show that a function chosen from a pairwise-independent family satisfies the property with high probability.

**Definition 1** Let  $A, B \subseteq \{0, 1\}^\ell$ ,  $h : \{0, 1\}^\ell \rightarrow \{0, 1\}^\ell$ , and  $\varepsilon > 0$ . We say  $h$  is  $(\varepsilon, A, B)$ -good if:

$$\left| \Pr_{x \in \{0, 1\}^\ell} \left[ x \in A \bigwedge h(x) \in B \right] - \Pr_{x, y \in \{0, 1\}^\ell} \left[ x \in A \bigwedge y \in B \right] \right| \leq \varepsilon.$$

**Lemma 3** Let  $A, B \subseteq \{0, 1\}^\ell$ ,  $H$  be a family of pairwise-independent functions, and  $\varepsilon > 0$ . Then:

$$\Pr_{h \in H} [h \text{ is not } (\varepsilon, A, B)\text{-good}] \leq \frac{\rho(A)\rho(B)}{2^\ell \varepsilon^2}.$$

**Proof** The proof is fairly straightforward. We want to bound the fraction of functions in  $H$  for which  $\left| \Pr_{x \in \{0,1\}^\ell} [x \in A \wedge h(x) \in B] - \rho(A) \cdot \rho(B) \right| > \varepsilon$  or, equivalently,

$$\left| \Pr_{x \in A} [h(x) \in B] - \rho(B) \right| > \varepsilon / \rho(A).$$

For fixed  $x$ , let  $\delta_{h(x) \in B}$  be an indicator random variable (over random choice of  $h \in H$ ) that is 1 iff  $h(x) \in B$ . Then we are interested in the fraction of  $h \in H$  for which

$$\left| \sum_{x \in A} \delta_{h(x) \in B} - |A| \cdot \rho(B) \right| > \varepsilon \cdot |A| / \rho(A).$$

Using Chebyshev's inequality and pairwise independence of  $H$ , we obtain

$$\Pr_{h \in H} \left[ \left| \sum_{x \in A} \delta_{h(x) \in B} - |A| \cdot \rho(B) \right| > \varepsilon \cdot |A| / \rho(A) \right] \leq \frac{|A| \cdot \rho(B) \cdot \rho(A)^2}{\varepsilon^2 |A|^2} = \frac{\rho(A) \rho(B)}{2^\ell \varepsilon^2}.$$

■

## 2.2 The Pseudorandom Generator and Its Analysis

### 2.2.1 The Basic Step

We first show how to reduce the number of random bits by roughly half. Let  $H$  denote a pairwise-independent family of functions, and fix an input  $x$ . Let  $Q$  denote the transition matrix corresponding to transitions in  $M_x$  after reading  $\ell$  random bits; that is, the  $(j, i)$ th entry of  $Q$  is the probability that  $M_x$ , starting in configuration  $i$ , moves to configuration  $j$  after reading  $\ell$  random bits. So  $Q^2$  is a transition matrix denoting the probability that  $M_x$ , starting in configuration  $i$ , moves to configuration  $j$  after reading  $2\ell$  random bits. Fixing  $h \in H$ , let  $Q_h$  be a transition matrix where the  $(j, i)$ th entry in  $Q_h$  is the probability that  $M_x$ , starting in configuration  $i$ , moves to configuration  $j$  after reading the  $2\ell$  “random bits”  $r \parallel h(r)$  (where  $r \in \{0, 1\}^\ell$  is chosen uniformly at random). Put differently,  $Q^2$  corresponds to taking two uniform and independent steps of a random walk, whereas  $Q_h$  corresponds to taking two steps of a random walk where the first step (given by  $r$ ) is random and the second step (namely,  $h(r)$ ) is a deterministic function of the first. We now show that these two transition matrices are “very close”. Specifically:

**Definition 2** Let  $Q, Q_h, \ell$  be as defined above, and  $\varepsilon \geq 0$ . We say  $h \in H$  is  $\varepsilon$ -good for  $Q$  if

$$\text{SD}(Q_h, Q^2) \leq \varepsilon / 2.$$

**Lemma 4** Let  $H$  be a pairwise-independent function family, and let  $Q$  be an  $N \times N$  transition matrix where transitions correspond to reading  $\ell$  random bits. For any  $\varepsilon > 0$  we have:

$$\Pr_{h \in H} [h \text{ is not } \varepsilon\text{-good for } Q] \leq \frac{N^6}{\varepsilon^2 2^\ell}.$$

**Proof** For  $i, j \in [N]$  (corresponding to configurations in  $M_x$ ), define

$$B_{i,j} \stackrel{\text{def}}{=} \{r \in \{0,1\}^\ell \mid r \text{ defines a transition from } i \text{ to } j\}.$$

For any fixed  $i, j, k$ , we know from Lemma 3 that the probability that  $h$  is not  $(\varepsilon/N^2, B_{i,j}, B_{j,k})$ -good is at most  $N^4 \rho(B_{i,j})\rho(B_{j,k})/\varepsilon^2 2^\ell \leq N^4 \rho(B_{i,j})/\varepsilon^2 2^\ell$ . Applying a union bound over all  $N^3$  triples  $i, j, k \in [N]$ , and noting that for any  $i$  we have  $\sum_j \rho(B_{i,j}) = 1$ , we have that  $h$  is  $(\varepsilon/N^2, B_{i,j}, B_{j,k})$ -good for *all*  $i, j, k$  except with probability at most  $N^6/\varepsilon^2 2^\ell$ .

We show that whenever  $h$  is  $(\varepsilon/N^2, B_{i,j}, B_{j,k})$ -good for all  $i, j, k$ , then  $h$  is  $\varepsilon$ -good for  $Q$ . Consider the  $(k, i)$ th entry in  $Q_h$ ; this is given by:  $\sum_{j \in [N]} \Pr_{r \in \{0,1\}^\ell} [r \in B_{i,j} \wedge h(r) \in B_{j,k}]$ . On the other hand, the  $(k, i)$ th entry in  $Q^2$  is:  $\sum_{j \in [N]} \rho(B_{i,j}) \cdot \rho(B_{j,k})$ . Since  $h$  is  $(\varepsilon/N^2, B_{i,j}, B_{j,k})$ -good for every  $i, j, k$ , the absolute value of their difference is

$$\begin{aligned} & \left| \sum_{j \in [N]} \left( \Pr[r \in B_{i,j} \wedge h(r) \in B_{j,k}] - \rho(B_{i,j}) \cdot \rho(B_{j,k}) \right) \right| \\ & \leq \sum_{j \in [N]} \left| \Pr[r \in B_{i,j} \wedge h(r) \in B_{j,k}] - \rho(B_{i,j}) \cdot \rho(B_{j,k}) \right| \\ & \leq \sum_{j \in [N]} \varepsilon/N^2 = \varepsilon/N. \end{aligned}$$

It follows that  $\text{SD}(Q_h, Q^2) \leq \varepsilon/2$  as desired. ■

The lemma above gives us a pseudorandom generator that reduces the required randomness by (roughly) half. Specifically, define a pseudorandom generator  $G_1 : \{0,1\}^{2\ell+R/2} \rightarrow \{0,1\}^R$  via:

$$G_1(h; r_1, \dots, r_{R/2\ell}) = r_1 \parallel h(r_1) \parallel \dots \parallel r_{R/2\ell} \parallel h(r_{R/2\ell}), \quad (1)$$

where  $h \in H$  (so  $|h| = 2\ell$ ) and  $r_i \in \{0,1\}^\ell$ . Assume  $h$  is  $\varepsilon$ -good for  $Q$ . Running  $M_x$  using the output of  $G_1(h; \dots)$  as the “random tape” generates the probability distribution

$$\overbrace{Q_h \cdots Q_h}^{R/2\ell} \cdot \mathbf{s}$$

for the final configuration, where  $\mathbf{s}$  denotes the initial configuration of  $M_x$  (i.e.,  $\mathbf{s}$  is the probability distribution that places probability 1 on the initial configuration of  $M_x$ , and 0 elsewhere). Running  $M_x$  on a truly random tape generates the probability distribution

$$\overbrace{Q^2 \cdots Q^2}^{R/2\ell} \cdot \mathbf{s}$$

for the final configuration. Since  $\text{SD}(Q_h, Q^2) \leq \varepsilon/2$ , we have

$$\text{SD}\left(\overbrace{Q_h \cdots Q_h}^{R/2\ell} \cdot \mathbf{s}, \overbrace{Q^2 \cdots Q^2}^{R/2\ell} \cdot \mathbf{s}\right) \leq \frac{R}{2\ell} \cdot \frac{\varepsilon}{2}.$$

This means that the behavior of  $M_x$  when run using the output of the pseudorandom generator is very close to the behavior of  $M_x$  when run using a truly random tape: in particular, if  $x \notin L$  then  $M_x$  in the former case accepts with probability at most

$$\Pr[\text{accepts} \wedge h \text{ is } \varepsilon\text{-good for } Q] + \Pr[h \text{ is not } \varepsilon\text{-good for } Q] \leq (1/4 + R\varepsilon/4\ell) + N^6/\varepsilon^2 2^\ell;$$

similarly, if  $x \in L$  then  $M_x$  in the former case accepts with probability at least  $3/4 - R\varepsilon/4\ell - N^6/\varepsilon^2 2^\ell$ . Summarizing (and slightly generalizing):

**Corollary 5** *Let  $H$  be a pairwise-independent function family, let  $Q$  be an  $N \times N$  transition matrix where transitions correspond to reading  $\ell$  random bits, let  $k > 0$  be an integer, and let  $\varepsilon > 0$ . Then except with probability at most  $N^6/\varepsilon^2 2^\ell$  over choice of  $h \in H$  we have:*

$$\text{SD} \left( \overbrace{Q_h \cdots Q_h}^k, \overbrace{Q^2 \cdots Q^2}^k \right) \leq k\varepsilon/2.$$

### 2.2.2 Recursing

Fixing  $h_1 \in H$ , note that  $Q_{h_1}$  is a transition matrix and so we can apply Corollary 5 to it as well. Moreover, if  $Q$  uses  $R$  random bits then  $Q_{h_1}$  uses  $R/2$  random bits (treating  $h_1$  as fixed). Continuing in this way for  $I \stackrel{\text{def}}{=} \mathcal{O}(\log(R/\ell))$  iterations, we obtain a transition matrix  $Q_{h_1, \dots, h_I}$ . Say all  $h_i$  are  $\varepsilon$ -good if  $h_1$  is  $\varepsilon$ -good for  $Q$ , and for each  $i > 1$  it holds that  $h_i$  is  $\varepsilon$ -good for  $Q_{h_1, \dots, h_{i-1}}$ . By Corollary 5 we have:

- All  $h_i$  are  $\varepsilon$ -good except with probability at most  $N^6 I / \varepsilon^2 2^\ell$ .
- If all  $h_i$  are  $\varepsilon$ -good then

$$\text{SD}(Q_{h_1, \dots, h_I}, \overbrace{Q^2 \cdots Q^2}^{R/2\ell}) \leq \frac{\varepsilon}{2} \cdot \sum_{i=1}^I \frac{R}{2^i \ell} = \mathcal{O}(\varepsilon R / \ell).$$

Equivalently, we obtain a pseudorandom generator

$$G_I(h_1, \dots, h_I; r) \stackrel{\text{def}}{=} G_{I-1}(h_1, \dots, h_{I-1}; r) \parallel G_{I-1}(h_1, \dots, h_{I-1}; h_I(r)),$$

where  $G_1$  is as in Equation (1).

### 2.2.3 Putting it All Together

We now easily obtain the desired derandomization. Recall  $N = 2^{\mathcal{O}(s)}$ . Set  $\varepsilon = 2^{-S}/10$ , and set  $\ell = \Theta(S)$  so that  $\frac{N^6 S}{\varepsilon^2 2^\ell} \leq 1/20$ . Then the number of random bits used (as input to  $G_I$  from the previous section) is  $\mathcal{O}(\ell \cdot \log(R/\ell) + \ell) = \mathcal{O}(S \log R)$  and the space used is bounded by that as well (using the fact that each  $h \in H$  can be evaluated using space  $\mathcal{O}(\ell) = \mathcal{O}(S)$ ). All  $h_i$  are good except with probability at most  $N^6 \log(R/\ell) / \varepsilon^2 2^\ell \leq N^6 S / \varepsilon^2 2^\ell \leq 1/20$ ; assuming all  $h_i$  are good, the statistical difference between an execution of the original algorithm and the algorithm run with a pseudorandom tape is bounded by  $2^{-S}/20 \cdot R \leq 1/20$ . Theorem 1 follows easily.

### 2.3 $BPL \subseteq SC$

A deterministic algorithm using space  $\mathcal{O}(\log^2 n)$  might potentially run for  $2^{\mathcal{O}(\log^2 n)}$  steps; in fact, as described, the algorithm from the proof of Corollary 2 uses this much time. For the particular pseudorandom generator we have described, however, it is possible to do better. The key observation is that instead of just choosing the  $h_1, \dots, h_I$  at random and simply hoping that they are all  $\varepsilon$ -good, we will instead deterministically search for  $h_1, \dots, h_I$  which *are* each  $\varepsilon$ -good. This can be done in polynomial time (when  $S = \mathcal{O}(\log n)$ ) because: (1) for a given transition matrix  $Q_{h_1, \dots, h_{i-1}}$  and candidate  $h_i$ , it is possible to determine in polynomial time and polylogarithmic space whether  $h_i$  is  $\varepsilon$ -good for  $Q_{h_1, \dots, h_{i-1}}$  (this relies on the fact that the number of configurations  $N$  is polynomial in  $n$ ); (2) there are only a polynomial number of possibilities for each  $h_i$  (since  $\ell = \Theta(S) = \mathcal{O}(\log n)$ ).

Once we have found the good  $\{h_i\}$ , we then cycle through all possible choices of the seed  $r \in \{0, 1\}^\ell$  and take majority (as before). Since there are a polynomial number of possible seeds (again using the fact that  $\ell = \Theta(S) = \mathcal{O}(\log n)$ ), the algorithm as a whole runs in polynomial time.

(For completeness, we discuss the case of general  $S = \Omega(\log n)$  assuming  $R = 2^S$ . Checking whether a particular  $h_i$  is  $\varepsilon$ -good requires time  $2^{\mathcal{O}(S)}$ . There are  $2^{\mathcal{O}(S)}$  functions to search through at each stage, and  $\mathcal{O}(S)$  stages altogether. Finally, once we obtain the good  $\{h_i\}$  we must then enumerate through  $2^{\mathcal{O}(S)}$  seeds. The end result is that  $BSPACE(S) \subseteq \text{TIME}SPC(2^{\mathcal{O}(S)}, S^2)$ .)

## 3 Applications to Error Reduction

Interestingly, the same pseudorandom generator we have constructed can also be used for efficient error reduction. Before discussing this application, we briefly discuss error reduction in general. (For simplicity, we focus here on the case of error reduction for randomized algorithms with *one-sided* error; all results described here can be generalized for the case of two-sided error.)

For concreteness, say we have an algorithm  $A$  for some language  $L$  such that

$$\begin{aligned} x \in L &\Rightarrow \Pr[A(x) = 1] \geq 1/2 \\ x \notin L &\Rightarrow \Pr[A(x) = 1] = 0. \end{aligned}$$

Say  $A$  uses  $\ell$  random bits. (The time/space complexity of  $A$  is not relevant to this discussion.) A naïve approach to error reduction would be to run  $A$  on a given input  $k$  times using independent random tapes  $r_1, \dots, r_k$ , outputting 1 iff any of these runs returns 1. This uses  $k \cdot \ell$  random bits, requires running  $A$  for  $k$  times, and achieves error  $2^{-k}$ .

A different approach (due to Chor-Goldreich) is to let the  $\{r_i\}$  be *pairwise independent* rather than completely independent. That is, choose random  $h \in H$  and set  $r_i = h(i) \in \{0, 1\}^\ell$ ; then run  $A$  for  $k$  times using the random coins  $r_1, \dots, r_k$ . This uses  $\mathcal{O}(\ell)$  random bits (the only randomness is the choice of  $h$ ) and  $k$  executions of  $A$  as before, but only achieves error  $\mathcal{O}(1/k)$ . (The proof follows directly from Chebyshev's inequality.)

A better approach uses (a small modification of) the pseudorandom generator from the previous section. Define  $G_1(h_1; r) = r \parallel h_1(r)$  and, inductively,

$$G_I(h_1, \dots, h_I; r) \stackrel{\text{def}}{=} G_{I-1}(h_1, \dots, h_{I-1}; r) \parallel G_{I-1}(h_1, \dots, h_{I-1}; h_I(r)).$$

(The difference from before is that now the output length of  $G_I$  grows; specifically, the output length of  $G_I$  is  $(\{0, 1\}^\ell)^{2^I}$ .) Our algorithm will now be to run  $A$  on each of the  $k \stackrel{\text{def}}{=} 2^I$  strings

output by  $G_I$ ; we output 1, as before, iff  $A$  outputs 1 in one of those executions. Now we use  $\mathcal{O}(\ell \cdot \log k)$  random bits (and  $k$  executions of  $A$ ); the error is given by the following theorem.

**Theorem 6** *If  $x \in L$ , the probability that  $A$  always outputs 0 when run on the  $k = 2^I$  random strings output by  $G_I$  is at most  $2^{-k} + (\log k + 2) \cdot 2^{-\ell/3}$ .*

**Proof** Setting  $\varepsilon = 2^{-\ell/3}$ , Lemma 3 shows that for any  $A, B \subseteq \{0, 1\}^\ell$  we have

$$\Pr_{h \in H} [h \text{ is not } (\varepsilon, A, B)\text{-good}] \leq \varepsilon.$$

Thus, all  $h_i$  are  $(\varepsilon, A, B)$ -good (for any  $A, B$ ) except with probability at most  $\varepsilon \cdot \log k$ .

Assuming all  $h_i$  are  $(\varepsilon, A, B)$ -good, we prove by induction on  $I$  that the probability (over choice of  $r \in \{0, 1\}^\ell$ ) that  $A$  always outputs 0 when run on the output of  $G_I(h_1, \dots, h_I; r)$  is at most  $2^{-2^I} + 2\varepsilon$ . For  $I = 0$  this is immediate. We prove it holds for  $I$ , assuming it holds for  $I - 1$ .

Let

$$A = B = \{r \mid A \text{ always outputs 0 when run on the output of } G_{I-1}(h_1, \dots, h_{I-1}; r)\}.$$

By our inductive step,  $\rho(A) = \rho(B) \leq 2^{-2^{I-1}} + 2\varepsilon$ . Furthermore, the probability that  $A$  always outputs 0 when run on the output of  $G_I(h_1, \dots, h_I; r)$  is exactly the probability that  $r \in A$  and  $h_I(r) \in B$ . Since  $h_I$  is  $(\varepsilon, A, B)$ -good we have

$$\begin{aligned} \Pr_{r \in \{0, 1\}^\ell} [r \in A \wedge h_I(r) \in B] &\leq \Pr_{r, r' \in \{0, 1\}^\ell} [r \in A \wedge r' \in B] + \varepsilon \\ &\leq \left(2^{-2^{I-1}} + 2\varepsilon\right)^2 + \varepsilon \\ &\leq 2^{-2^I} + 2\varepsilon. \end{aligned}$$

This completes the proof. ■

## Bibliographic Notes

The results of Section 2 are due to [3, 4], both of which are very readable. See also [1, Lecture 16] for a slightly different presentation. Section 3 is adapted from the Luby-Wigderson survey on pairwise independence [2].

## References

- [1] O. Goldreich. Introduction to Complexity Theory (July 31, 1999).
- [2] M. Luby and A. Wigderson. *Pairwise Independence and Derandomization*. Foundations and Trends in Theoretical Computer Science. Now Publishers Inc., 2006. (Available freely online.)
- [3] N. Nisan. Pseudorandom Generators for Space-Bounded Computation. *STOC '90*.
- [4] N. Nisan.  $RL \subseteq SC$ . *Computational Complexity 4*: 1–11, 1994. (Preliminary version in *STOC '92*.)